

## 7.5.4 Conditional Variance

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$$\text{Var}(X|Y) = E[(X - E[X|Y])^2 | Y] \quad (\text{defn})$$

$$\text{Var}(X|Y) = E[X^2|Y] - (E[X|Y])^2 \quad (\text{alt. formula})$$

Theorem (see book for derivation):

$$\text{Var}(X) = E[\text{Var}(X|Y)] + \text{Var}(E[X|Y])$$

**Example**  
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Suppose that by any time  $t$  the number of people who have arrived at a train depot is a Poisson random variable with mean  $\lambda t$ . If the initial train arrives at the depot at a time (independent of when the passengers arrive) that is uniformly distributed over  $(0, T)$ , what are the mean and variance of the number of passengers who enter the train?

$N(t)$ : # of passengers that arrive by time  $t$ .

$N(t) \sim \text{Poisson}(\lambda t)$ ,  $\lambda > 0$  some fixed number

$Y$ : time when train arrives

$Y \sim \text{Unif}(0, T)$ ,  $T > 0$  some fixed number

$N$  and  $Y$  are indep.

$X$ : # of passengers who enter the train

$$X = N(Y)$$

(a)  $E[X] = ?$ , (b)  $\text{Var}(X) = ?$

The idea is to condition on  $Y$ .

$$(a) \quad E[X] = E[N(Y)]$$

$$= E[E[N(Y)|Y]]$$

$$E[N(Y)|Y=y] = E[N(y)] = \lambda y$$

b/c  $N$  and  $Y$   
are indep.

Poisson( $\lambda y$ )

$$\text{So... } E[N(Y)|Y] = \lambda Y$$

$$E[X] = E[E[N(Y)|Y]]$$

$$= E[\lambda Y]$$

$$= \lambda E[Y] = \lambda \cdot \frac{T}{2} = \boxed{\frac{\lambda T}{2}}$$

$$(b) \quad \text{Var}(X) = E[\text{Var}(X|Y)] + \text{Var}(E[X|Y])$$

$$\text{Var}(X|Y) = \text{Var}(N(Y)|Y)$$

$$\text{Var}(N(Y)|Y=y) = \text{Var}(N(y)) = \lambda y$$

b/c  $N$  and  $Y$   
are indep. Poisson( $\lambda y$ )

$$\text{So... } \text{Var}(N(Y)|Y) = \lambda Y \checkmark$$

$$E[X|Y] = E[N(Y)|Y] \stackrel{\substack{\uparrow \\ \text{from part (a)}}}{=} \lambda Y$$

$$\text{Var}(X) = E[\lambda Y] + \text{Var}(\lambda Y)$$

$$= \lambda E[Y] + \lambda^2 \text{Var}(Y)$$

$$= \lambda \cdot \frac{T}{2} + \lambda^2 \cdot \frac{T^2}{12}$$

$$= \frac{\lambda^2 T^2 + 6\lambda T}{12}$$

Notice that I didn't need any integrals or series. If you understand conditional exp. conceptually, you often don't.

Unif(0, T)

## 7.7 Moment generating functions

$$M_X(t) = E[e^{tX}] \quad \text{moment generating func. (MGF)}$$

domain of  $M_X$  is all  $t$ 's such that the expected value exists (it fails to exist when the series or integral is divergent)

Always exists for  $t=0$  and  $M_X(0) = 1$ .

Book derives

$$E[X^n] = M_X^{(n)}(0)$$

$n^{\text{th}}$  derivative of  $M_X$

the "moments" of  $X$

If  $X$  and  $Y$  are indep., then

$$M_{X+Y}(t) = M_X(t)M_Y(t)$$

check:  $M_{X+Y}(t) = E[e^{t(X+Y)}] = E[e^{tX+tY}]$   
 $= E[e^{tX}e^{tY}] = E[e^{tX}]E[e^{tY}] = M_X(t)M_Y(t)$   
↑ indep.

Look at and get familiar w/ the tables of common MGFs. (Book derives a few of them.)

Table 7.1, p. 364 for discrete

Table 7.2, p. 365 for continuous

Key thm: If  $X$  and  $Y$  have the same MGF, then they have the same distribution

∴ Can use MGF to figure out distribution.  
(compute MGF and look it up on the table)

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**Example**

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**Sums of independent Poisson random variables**

Calculate the distribution of  $X + Y$  when  $X$  and  $Y$  are independent Poisson random variables with means respectively  $\lambda_1$  and  $\lambda_2$ .

$$X \sim \text{Poisson}(\lambda_1), \quad Y \sim \text{Poisson}(\lambda_2)$$

$$X, Y \text{ indep.} \Rightarrow M_{X+Y}(t) = M_X(t) M_Y(t)$$

$$\left. \begin{aligned} M_X(t) &= e^{\lambda_1(e^t-1)} \\ M_Y(t) &= e^{\lambda_2(e^t-1)} \end{aligned} \right\} \text{ from the tables}$$

$$M_{X+Y}(t) = e^{\lambda_1(e^t-1)} e^{\lambda_2(e^t-1)} = e^{(\lambda_1 + \lambda_2)(e^t-1)}$$

From the table, we must have

$$X+Y \sim \boxed{\text{Poisson}(\lambda_1 + \lambda_2)}$$

**Example**  
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### Sums of independent normal random variables

Show that if  $X$  and  $Y$  are independent normal random variables with respective parameters  $(\mu_1, \sigma_1^2)$  and  $(\mu_2, \sigma_2^2)$ , then  $X + Y$  is normal with mean  $\mu_1 + \mu_2$  and variance  $\sigma_1^2 + \sigma_2^2$ .

$$X \sim N(\mu_1, \sigma_1^2), Y \sim N(\mu_2, \sigma_2^2)$$

$X, Y$  indep.

$$\left. \begin{aligned} M_X(t) &= e^{\mu_1 t + \sigma_1^2 t^2 / 2} \\ M_Y(t) &= e^{\mu_2 t + \sigma_2^2 t^2 / 2} \end{aligned} \right\} \text{ from the tables}$$

$$M_{X+Y}(t) = M_X(t) M_Y(t)$$

$$= e^{\mu_1 t + \sigma_1^2 t^2 / 2} e^{\mu_2 t + \sigma_2^2 t^2 / 2}$$

$$= e^{(\mu_1 + \mu_2)t + (\sigma_1^2 + \sigma_2^2)t^2/2}$$

From the table, we must have

$$X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

HW: Ch. 7 : Problems 70, 80