3.5 $P(1F)$ is a probability	83
The function $P(1F)$ satisfies the axioms of probability.	
2.2 $Q(n)$ The second, that's true for $P(1)$ is true for $P(1)$	
3.3 $P(n)$ the second, that's true for $P(1)$ is true for $P(1)$	
4.4 $P(n) = P(8)P(A 8) + P(8)P(A 8)$	
4.4 $P(n) = P(8)P(A 8) + P(8)P(A 8)P(A 8)$	
4.4 $P(n) = P(8)P(A 8) + P(8)P(A 8)P(A 8)$	
4.4 $P(n) = P(8)P(A 8) + P(8)P(A 8)$	
4.4 $P(n) = P(8)P(A 8)P(A 8)$	
5.4 $P(1)P(1)P(A 8)P(A 8)$	
6. $P(1)P(1)P(A 8)P(A 8)$	
7. $P(1)P(A 8)P(A 8)$	
8. $P(1)P(A 8)P(A 8)$	
9. $P(1)P(A 8)P(A 8)$	
10. $P(1)P(A 8)P(A 8)$	
11. $P(1)P(A 8)P(A 8)$	
12. $P(1)P(A 8)P(A 8)$	
13. $P(1)P(A 8)P(A 8)$	
14. $P(1)P(A 8)P(A 8)$	
15. $P(1)P(A 8)P(A 8)$	
16. $P(1$	

Example The gambler's ruin problem

 4_m

Two gamblers, A and B , bet on the outcomes of successive flips of a coin. On each flip, if the coin comes up heads, A collects 1 unit from B , whereas if it comes up tails, A pays 1 unit to B . They continue to do this until one of them runs out of money. If it is assumed that the successive flips of the coin are independent and each flip results in a head with probability p , what is the probability that A ends up with all the money if he starts with i units and B starts with $N - i$ units?

(This example is from Section 3.4.)

\n
$$
E = \n\begin{bmatrix}\n\mu_{0} & \mu_{0} & \mu_{0} & \mu_{0} \\
\mu_{i} &= \mu_{i} & \mu_{i} & \mu_{i} & \mu_{i}\n\end{bmatrix}
$$
\n
$$
P(E|E_{i}) = ?
$$
\n
$$
P_{i} = P(E|E_{i})
$$
\n
$$
P_{i} = 0
$$
\n
$$
P_{i} = ?
$$
\n
$$
P_{i} = P(E|E_{i}) = \text{Rearl}
$$
\n
$$
P_{i} = P(E|E_{i}) = \text{Rartiv}_{i} \quad \text{for } i \neq i \text{ and } i \neq j \text{
$$

$$
+PAE(E;)P(E|H^C A E))
$$

$$
\nabla^{\times} \theta^{\leq L} \quad P_{i} = \phi P_{i\star i} + g P_{i-1}
$$
\n
$$
(\phi + g) P_{i} = \phi P_{i\star i} + g P_{i-1}
$$
\n
$$
\phi P_{i} \rightarrow g P_{i} = \phi P_{i\star i} + g P_{i-1}
$$
\n
$$
g(P_{i} - P_{i-1}) = \phi (P_{i+1} - P_{i}) \qquad \text{where } g \text{ is the }
$$
\n
$$
P_{i\star i} - P_{i} = \frac{g}{\phi} (P_{i} - P_{i-1})
$$

$$
P_0 = 0
$$

\n $P_2 - P_1 = \frac{q}{p} (P_1 - \hat{P_0}) = \frac{q}{p} P_1$
\n $P_3 - P_2 = \frac{q}{p} (P_2 - P_1) = (\frac{q}{p})^2 P_1$

$$
P_{N} - P_{N-1} = \left(\frac{q}{p}\right)^{N-1} P_{1}
$$

\n
$$
S_{0} ... \text{ for any } k, \text{ sum}
$$
\n
$$
P_{k} = P_{k} - P_{0} = \sum_{j=1}^{k} (P_{j} - P_{j-1}) = \sum_{j=1}^{k} \left(\frac{q}{p}\right)^{j-1} P_{1}
$$

$$
\mathsf{v}:=\frac{\mathcal{B}}{\mathcal{P}}
$$

 $\mathcal{L}^{\text{max}}_{\text{max}}$ and $\mathcal{L}^{\text{max}}_{\text{max}}$

$$
P_k = P_1 \sum_{j=1}^{k} r^{j-1}
$$

but $P_i = ?$ (with need to use $P_N = 1$
to figure that out)

Case 1 :
$$
p \neq \frac{1}{2}
$$

\nIn this case, $r \neq 1$, so
\n $P_k = P_1 \sum_{j=1}^{k} r^{j-1} = P_1 \cdot \frac{1-r^{k}}{1-r}$
\n P_{α} from 1
\n P_{α} from 2
\n P_{α} from 3
\n P_{α} from 4
\n $P_{\alpha} = P_1 \cdot \frac{1-r^{N}}{1-r} = 1$, so...
\n $\frac{1}{2} = P_1 \cdot \frac{1-r^{N}}{1-r} = \frac{1-r^{N}}{1-r^{N}}$
\n $\therefore P_k = \frac{1-r}{1-r^{N}} \cdot \frac{1-r^{k}}{1-r} = \frac{1-r^{k}}{1-r^{N}}$
\nSub in $r = \frac{q}{p}$:
\n $P_k = \frac{1-(\frac{q}{p})^{k}}{1-(\frac{q}{p})^{k}} \cdot \frac{p^{N}}{p^{N}} = \frac{p^{N}-p^{N-k}}{p^{N}-q^{N}}$

Case 2:
$$
p = \frac{1}{2}
$$

\n l_{w} this case, $r=1$, so
\n $P_{k} = P_{1} \sum_{j=1}^{k} r^{j-1} = P_{1} k$
\nWe know that $P_{N} = 1$, so...
\n $1 = P_{1} N \Rightarrow P_{1} = \frac{1}{N}$
\n $\therefore P_{k} = \frac{k}{N}$

Final answer

$$
P_{k}=P(E|E_{k})=\left\{\begin{array}{c}\frac{p^{N}-p^{N-k}q^{k}}{N} & \text{if } p\neq\frac{1}{2},\\ \frac{k}{N} & \text{if } p=\frac{1}{2}.\end{array}\right\}
$$

Expl 4m continued Let $F = "A goes broke."$ What is the relationship between E and F ? It's tempting to think they're negations of one another. But they're not! $F \times E$ The reason is because there's a theird possibility: G = "The game goes on forever lie. no one ever wins all the money E E , F , and G form a partition (exactly one of them is true). $P(G | E_i) = ?$ $P(E | E_i) + P(F | E_i) + P(G | E_i) = 1$ we know this re know this if we can find this, we
already can solve for P(GIE; can solve for P(GIE;) $P(G | E_i) = 1 - (P(E | E_i) + P(F | E_i))$

Let F: = "B starts with \$i." By symmetry, $P(F|F_i)$ is the same as P(E(Ei), but with the roles of p and g reversed:

$$
A1_{S_{v}}, F_{i} = E_{N-i} \cdot S_{o...}
$$
\n
$$
P(F|E_{i}) = P(F|F_{N-i})
$$
\n
$$
= \begin{cases}\n\frac{a^{N} - a^{i}p^{N-i}}{a^{N} - p^{N}} & \text{if } p \neq \frac{1}{2}, \\
\frac{N-i}{N} & \text{if } p = \frac{1}{2}.\n\end{cases}
$$

Nou ve can compute P(GIE):

Case 1:
$$
P = \frac{1}{2}
$$

\n $P(G | E_i) = 1 - (P(E | E_i) + P(F | E_i))$
\n $= 1 - (\frac{i}{N} + \frac{N - i}{N})$
\n $= 1 - 1$
\n $= 0$
\nCase 2: $p \neq \frac{1}{2}$
\n $P(G | E_i) = 1 - (P(E | E_i) + P(F | E_i))$
\n $= 1 - (\frac{p^N - p^{N-i}q^i}{p^N - q^N} + \frac{q^{N} - q^i p^{N-i}}{q^{N} - p^N})$
\n $= 1 - (\frac{p^N - p^{N-i}q^i - q^N + q^i p^{N-i}}{p^N - q^N})$

 $= 1 - 1$ $= 0$. In either case, $\sqrt{P(G|E_i)}=0$

Laplace's rule of succession **Example**

5e

There are $k + 1$ coins in a box. When flipped, the *i*th coin will turn up heads with probability i/k , $i = 0, 1, ..., k$. A coin is randomly selected from the box and is then repeatedly flipped. If the first n flips all result in heads, what is the conditional probability that the $(n + 1)$ flip will do likewise?

$$
P(A_{n+1} | A) = \sum_{i=0}^{k} P(C_i | A) P(A_{n+1} | C_i \cap A)
$$
\n
$$
= \sum_{i=0}^{k} P(C_i | A) P(A_{n+1} | C_i)
$$
\n
$$
= \sum_{i=0}^{k} P(C_i | A) P(A_{n+1} | C_i)
$$

$$
P(C_i | A) = \frac{P(C_i) P(A | C_i)}{\sum_{j=0}^{k} P(C_j) P(A | C_j)}
$$

$$
P(A|C_{j}) = P(A, A_{2} \cap \cdots \cap A_{n}(C_{j})
$$
\n
$$
\frac{C}{n} \cdot \frac{C}{n} P(A, C_{j}) P(A_{2}(C_{j}) \cdots P(A_{n}(C_{j}))
$$
\n
$$
= (\frac{j}{k})^{n}
$$
\n
$$
P(4, C_{j}) P(A_{2}(C_{j}) \cdots P(A_{n}(C_{j}))
$$
\n
$$
= (\frac{j}{k})^{n}
$$
\n
$$
P(4, C_{j}) P(A_{2}(C_{j}) \cdots P(A_{n}(C_{j}))
$$

$$
P(C;|A) = \frac{\frac{1}{k+1}(\frac{1}{k})}{\sum_{j=0}^{k} \frac{1}{k+1}(\frac{j}{k})^{n}}
$$
 conjugate

$$
P(A_{n+1}|A) = \sum_{i=0}^{k} P(C_{i}(A) P(A_{n+1} | C_{i})
$$
\n
$$
= \sum_{i=0}^{k} \left(\frac{\frac{1}{k+1} (\frac{i}{k})^{n}}{\frac{1}{j+0} \cdot \frac{1}{k+1} (\frac{i}{k})^{n}} \right) \cdot \frac{i}{k}
$$
\n
$$
= \frac{\left(\sum_{i=0}^{k} \frac{1}{k+1} (\frac{i}{k})^{n+i} \right)}{\left(\sum_{j=0}^{k} \frac{1}{k+1} (\frac{j}{k})^{n}} \right)}
$$
\nThe above sums are Riemann sums.
\nSo if k is long.
\n
$$
P(A_{n+1}|A) \approx \frac{\int_{0}^{1} x^{n+1} dx}{\int_{0}^{1} x^{n} dx} = \frac{\frac{1}{n+2} x^{n+1} b}{\frac{1}{n+1} x^{n+1} b}
$$
\n
$$
P(A_{n+1}|A) \approx \frac{\int_{0}^{1} x^{n} dx}{\int_{0}^{1} x^{n} dx} = \frac{\frac{1}{n+2} x^{n+1} b}{\frac{1}{n+1} x^{n+1} b}
$$